

# ECON4135: Solution to Written Paper 2

25th October 2007

## General comments

Some general comments regarding the problem set, and your answers:

- Most answers were not very good. You'll need to improve, in order to get good grades on the exam. Read through this solution set to see what was expected (some tips and comments are also included, these were not expected).
- Some of you misunderstood some of the problems. I've generally been fairly tolerant and employed considerable good will when correcting.
- You should always include output from Stata, to show what you have done (preferably logs stating both commands used, and the results you get - see some more comments in footnote 1). However, if you have much Stata-output, it may be better to append it to the paper, rather than include it in the text, as this make the paper difficult to read.
- Remember, you are *economists*, not statisticians! So, while it of course is crucial to do the estimation and calculations correctly, don't stop there. Try to give your results a (brief) economic interpretation. This is of course particularly important when you are explicitly asked to comment or interpret the results.
- Also, proofread your writing. Try to avoid nonsensical sentences, and generally try to be concise.

## Problem 1

See the appended log-file for Stata-commands used and the output Stata produced.<sup>1</sup> <sup>2</sup> I have summarized the 2003 data, results for 2004 are similar. From the log we see that we have 4215 firms, but with some missing observations on the *VA\_empl*-variable. Number of employees

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<sup>1</sup>The appended log documents the entire Stata-session, answering all problems of this paper. When reporting results, you should always include a (part of a) log, specifying both the command you used, and the output from the program. You can do this either by inserting the output into the document, or by appending a log. When estimating large models you may exclude the irrelevant coefficients. The appended or inserted log will be a lot easier to read if you use a fixed width font (for example *courier*) or format the regression output as a table.

<sup>2</sup>A note on Stata syntax: All commands, and all variables in Stata, can be abbreviated, as long as they are still unambiguous. Thus, writing `sum RD` is equivalent to writing `summarize RD_subsidy`. Also, when simultaneously referring to several variables, it may not be necessary to write all their names, see `help varlist` in Stata.

ranges from 0 (possibly misreporting?) to 3378, with an average of 30. Average tax deduction (for all firms) was 59' kroner, but only 14 percent of the firms actually got a deduction.

The firms which got a deduction are, on average, larger (mean number of employees is 61), have more highly educated employees and have a higher value added, but a larger share of these firms have no payable tax. The average subsidy within those firms who did get something is 420' kroner, ranging from 2.5' kroner to 1.6M kroner (which, interestingly, is exactly twice the upper limit). The median (at 352' kroner) is smaller than the mean, and the distribution seems to be skewed to the right. This is not unexpected, given that you cannot get a subsidy smaller than 0, but some few firms will get large subsidies.

## Problem 2

We want to estimate the equation

$$\begin{aligned} \ln RDsubsidy_i = & \beta_0 + \beta_1 \cdot taxposition_i + \beta_2 \cdot share\_high_i \\ & + \beta_3 \cdot VA\_empl_i + \beta_4 \cdot firmage\_10y_i + \beta_5 \cdot employ + u_i \end{aligned} \quad (1)$$

In order to estimate it, we make the following assumptions about the error term,  $u_i$ :

$$E(u_i|X_i) = 0 \quad (2)$$

$$(RDsubsidy_i, X_i) \text{ are } i.i.d. \text{ vectors} \quad (3)$$

$$var(u_i|X_i) = \sigma^2 \quad (4)$$

In the equations above  $X_i$  refers to the entire vector of covariates, i.e. *taxposition*, *share\_high* etc.

Assumption (2) is essential, it ensures that  $cov(u_i, X_i) = 0$ . This is required for the OLS-estimators to be unbiased and consistent, that is the estimators are on average correct, and as the number of observations increases the probability that the estimators will be very different from the true values becomes small.<sup>3</sup>

Assumption (3) assures that the error terms are independent across observations.

The last assumption is of *homoskedasticity* (i.e., equal error term variance across all observations). This assumption is *not* required for the OLS-estimators to be unbiased or consistent, but if it is not satisfied the estimated standard errors of the OLS-estimators will be misleading, and there will exist other unbiased estimators with smaller standard errors. We'll soon return to this.

Using Stata to estimate the model for 2003, we get the results shown below. You can find results for 2004 in the appended log.

```
. reg RD tax share VA firm employ
```

Source	SS	df	MS	
Model	5176058.31	5	1035211.66	Number of obs = 4084
Residual	133374640	4078	32705.8951	F( 5, 4078) = 31.65
				Prob > F = 0.0000
				R-squared = 0.0374
				Adj R-squared = 0.0362

<sup>3</sup>Note that as long as we include a constant term in the regression,  $E(u_i) = 0$  is not restrictive. The critical part of assumption (2) is that the covariates does not contain any information about the error term.

Total | 138550699 4083 33933.5534                      Root MSE           = 180.85

RD_subsidy	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
taxposition	-25.62379	5.72527	-4.48	0.000	-36.84845	-14.39914
share_high	220.8335	25.34942	8.71	0.000	171.1348	270.5322
VA_empl	.0027261	.0037569	0.73	0.468	-.0046394	.0100916
firmage_10y	-8.516276	5.701123	-1.49	0.135	-19.69359	2.661037
emply	.1238319	.0173638	7.13	0.000	.0897894	.1578744
_cons	66.48483	5.233204	12.70	0.000	56.2249	76.74477

From the computer output we see that the conditional expectation is given as<sup>4</sup>

$$E(RDsubsidy_i|X_i) = 66.5 - 25.6 \cdot taxposition_i + 221 \cdot share\_high_i + .00273 \cdot VA\_empl_i - 8.52 \cdot firmage\_10y_i + .124 \cdot emply$$

Thus, we see that there is a negative correlation between subsidy and positive payable tax, while the share of highly educated employees and number of employees both correlates positively with the subsidy. There is no significant effect<sup>5</sup> of value added or age of the firm.

The assumption of homoskedasticity may be overly restrictive (optimistic?), heteroskedasticity is often a problem in cross-sectional data like these. This means that the error term variance may not be constant over firms ( $var(u_i|S_i, E_i) = \sigma_i^2$ ), for example we expect the range of potential variation to be larger for larger firms (e.g. firms with more employees). In order to handle this we can use robust standard errors, as is done in the regression output below (still using data for 2003, with results for 2004 given in the appendix):

```
. reg RD tax share VA firm emply ,robust
```

Linear regression	Number of
obs = 4084	
	F( 5, 4078) = 12.15
	Prob > F = 0.0000
	R-squared = 0.0374
	Root MSE = 180.85

RD_subsidy	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
taxposition	-25.62379	5.712823	-4.49	0.000	-36.82404	-14.42354

<sup>4</sup>When reporting regression results, try to use a meaningful level of precision. Report at least the two first non-zero digits, if you report  $\beta_3 = 0.003$ , this could mean anything from 0.0025 to 0.0035, which may be an important difference. However, it is seldom relevant to report more than three to four digits either, the twelfth digit will typically neither be precisely estimated nor interesting. However, when doing calculations you should include a few extra decimal places to avoid error due to lacking numerical precision.

<sup>5</sup>If you are to be prudent, 'effect' is a strong word. It implies a statement about causality, which may not always be warranted.

share_high		220.8335	41.06181	5.38	0.000	140.33	301.3371
VA_empl		.0027261	.0053747	0.51	0.612	-.0078113	.0132635
firmage_10y		-8.516276	5.637936	-1.51	0.131	-19.56971	2.537156
empl		.1238319	.0444131	2.79	0.005	.0367579	.2109059
_cons		66.48483	5.617412	11.84	0.000	55.47164	77.49803

Much is unchanged: All the coefficient estimates and the  $R^2$ . The `robust`-option makes Stata calculate the estimated standard errors in a different way however, so these are different, and thus the  $t$ - and  $p$ -values also change. More specifically, the standard errors increase (except for *firmage\_10y* and *taxposition*, which are marginally reduced), but to a different degree: While the standard errors of *share\_high* and *VA\_empl* increase somewhat, there is a more than twofold increase in the standard errors of *empl*.

### Problem 3

In order to increase the fit of the model, we want to replace  $\beta_5 \cdot \text{empl}$  in eq. (1) with the dummy set  $\sum_{k=2}^5 \gamma_k \cdot \text{empl}\{k\}$ . The regression output is given below.<sup>6</sup>

```
. reg RD tax share VA firm empl2-empl5
```

Source	SS	df	MS	Number of obs =	4084
Model	16025704.5	8	2003213.06	F( 8, 4075) =	66.62
Residual	122524994	4075	30067.4832	Prob > F =	0.0000
Total	138550699	4083	33933.5534	R-squared =	0.1157
				Adj R-squared =	0.1139
				Root MSE =	173.4

RD_subsidy	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
taxposition	-20.63591	5.498826	-3.75	0.000	-31.41661 -9.855205
share_high	229.1315	24.36033	9.41	0.000	181.3719 276.891
VA_empl	-.0000786	.0036064	-0.02	0.983	-.007149 .0069918
firmage_10y	7.31752	5.552952	1.32	0.188	-3.569299 18.20434
empl2	22.09559	7.289712	3.03	0.002	7.803771 36.38741
empl3	75.84837	6.724101	11.28	0.000	62.66546 89.03128
empl4	187.8132	13.00368	14.44	0.000	162.3189 213.3076
empl5	189.2877	13.37939	14.15	0.000	163.0568 215.5186
_cons	14.61535	6.425678	2.27	0.023	2.017514 27.21319

We see that this model fits better,  $R^2$  increases from 0.037 to 0.116. Also, the similar increase in adjusted  $R^2$  indicates that this reflects a true increase in explanative power, not just more

<sup>6</sup>I have, for the added output, chosen to stick to non-robust estimation in the following, even though the results in the last problem indicated we may have an issue with heteroskedasticity here. However, remember that even if this is the case, our estimates are still unbiased and consistent, but the standard errors may be misleading.

variables. Inspecting the coefficients, we see find the cause: they do not reflect a linear relationship. Rather, after a rapid initial rise in the expected subsidy with employees, the marginal effect of additional employees becomes practically zero.<sup>7</sup> In light of this non-linearity, I'll stick to the dummy specification for the rest of the problem set.

If we also included *empl1* in the model, we would have the situation that  $\sum_{k=1}^5 \text{empl}\{k\} = 1 = \text{constant term}$ . I.e., some of the variables in the regression would be a linear combination of each other, and we would have a problem with multicollinearity. This makes it impossible to estimate the model, we cannot distinguish between the effect of the employment categories and the constant term. In order to avoid this problem, we must eliminate one variable, to make this the reference category (Stata would automatically have dropped one of the employment-dummies).<sup>8</sup>

Perfect multicollinearity is rarely a problem with continuous variables, these are very unlikely to be linear combinations of each other<sup>9</sup>, but when using categorical variables it is easy to include a complete set of dummies.

There is nothing special about *empl1*, so we could just as well have excluded any of the other categories to make up the reference (or we could have excluded the constant term). This will change the estimated employment-coefficients, but the standard errors will not change, and neither will the differences between the coefficients. Thus, if we rather excluded *empl2* (this amount to forcing  $\hat{\gamma}_2 = 0$ ), we would get  $\hat{\gamma}_1 = -22.1$ ,  $\hat{\gamma}_3 = 75.8 - 22.1 = 53.7$  etc. As the coefficient changes, and the standard errors stay the same, the *t*- and *p*-values will also change. This is because the coefficients are now tested against a different null hypothesis.

## Problem 4

A 99% CI is given as

$$[\hat{\beta} - t_{df,0.005}^c \cdot \hat{se}, \hat{\beta} + t_{df,0.005}^c \cdot \hat{se}], \quad (5)$$

where  $t_{df,0.005}^c$  is the critical *t*-value for a two-sided test at the 99% level of significance, using a *t*-distribution with  $df = N - K - 1$  (the number of observations minus the number of covariates minus one (for the constant term)) degrees of freedom. Below I show how to calculate the lower bound for the coefficient of *VA\_empl*, using Stata's `display`-command, and the stored values from the estimation:<sup>10</sup>

```
. di "VA_empl, 99% ci lower bound: " %9.5g _b[VA] - invttail(e(df_r),0.005)*_se[VA]
VA_empl, 99%ci lower bound: -.0093723
```

<sup>7</sup>This indicates that a promising, and more parsimonious specification could be to rather use some concave function of *empl*, such as  $\log(\text{empl})$  - try it!

<sup>8</sup>The problem set explicitly asks you to answer this, without reestimating the model. Then you should do just that, almost all of you have estimated the model with the extra dummy. If this is necessary for you to see what will happen, you will need to study for the exam!

<sup>9</sup>It is possible that such variables are highly, but not perfectly correlated, this is called imperfect multicollinearity. In such cases estimation is possible, but standard errors increase.

<sup>10</sup>Stata stores coefficients in the vector `_b`, standard errors in the vector `_se`, and several other useful stuff in `e(.)` - try running `ereturn list` after an estimation. Using these saves copying, marginally increases precision, and is extremely practical if you want to write programs (as opposed to using Stata interactively).

An easier way of doing this, however, is just to get the 99% CI's directly from the estimation, using the option `level(99)`.<sup>11</sup> See the appended log for Stata command and output.<sup>12</sup> Thus we directly get the relevant CI's:

```
share_high : [166, 292]
VA_empl   : [-.00937, .00921]
firmage_10y : [-6.99, 21.6]
```

The meaning of a 99% CI is that if we estimate a (correctly specified!) regression model on many (independent) samples, the CI would encompass the true parameter value 99% of the times.

We see that the confidence intervals for both *VA\_empl* and *firmage\_10y* encompass zero, thus we conclude that of these three variables, only *share\_high* is a significant determinant of the subsidy. Firms with a higher share of employees at the highest educational level tend to get a larger subsidy. This may reflect that most R&D is done by highly educated staff, thus the expected amount of R&D done in a firm, and thus the subsidy, increases with this share.

## Problem 5

From the above regression output, and the one for 2004 in the appended log, we see that 99% CI's for *taxposition* are:<sup>13</sup>

```
2003 : [-34.8, -6.47]
2004 : [-51.2, -21.5]
```

We see that although the coefficients may seem to be different, the confidence intervals do overlap, so we have no strong evidence for claiming there is a change. A test statistic to check for this could be:

$$\begin{aligned}
 t &= \frac{\hat{\beta}_{2003} - \hat{\beta}_{2004}}{\hat{se}(\hat{\beta}_{2003} - \hat{\beta}_{2004})} = \frac{\hat{\beta}_{2003} - \hat{\beta}_{2004}}{\sqrt{\hat{se}(\hat{\beta}_{2003})^2 + \hat{se}(\hat{\beta}_{2004})^2}} \\
 &= \frac{-20.63591 - -36.38267}{\sqrt{5.498826^2 + 5.768243^2}} = 1.98,
 \end{aligned}$$

see the appended log for calculations.<sup>14</sup> This is to be compared with a critical value, for a two-sided test with 99% level of significance, the value of 2.58 from the normal distribution gives a more than sufficient approximation here. We see that  $t < t^c$ , and thus conclude, as we did from inspecting the CI's, that we can not reject the null hypothesis, of unchanged coefficients.

---

<sup>11</sup>Yet another option, if you're just interested in one or a few variables, and don't want to run the entire regression (this may be a hassle, if the number of observations and covariates is large) is Stata's command `lincom`. See the log-file, and look it up in Stata's help system!

<sup>12</sup>Some of you estimated a regression equation containing just the variables for which you need a CI. You should rather stick to the complete specification, controlling for other covariates as well.

<sup>13</sup>Very many of you estimated a regression equation containing just *taxposition*. You should rather stick to the complete specification, controlling for other covariates as well.

<sup>14</sup>Testing is a more of a hassle in this case, because I'm using results from two different regressions, and thus can't use Stata's internal commands. Thus, the test in Problem 7 is easier to perform. For this test, I've just copied the values from the regression output.

## Problem 6

See the appended log for the relevant commands and Stata output. From the log we see that the estimated coefficient of the 2003 dummy is -8.40, and that this is significant at the 95% level of significance (although not at the 99% level). Thus, expected subsidies are larger in 2004 than in 2003. This may reflect several different explanations, e.g. the subsidies may be adjusted to reflect inflation, although in that case the increase may seem large (compare to the average of 58.7 found for 2003 in problem 2004). Other possible explanations may be increased funding for the scheme resulting in larger pay-outs, the firms may have increased research or just gotten better at writing applications.

If we included a dummy also for 2004, we would have that  $d_{2003} + d_{2004} = 1 = \text{constant}$ , i.e. multicollinearity, as in Problem 3. Thus, in order to be able to estimate the model, Stata would have dropped either of the dummies.

## Problem 7

The relevant test of the relationship between the expected subsidy and *taxposition* is just the regression in the last problem. Thus, from the Stata output associated with Problem 6, we see that *taxposition* is negatively related to the subsidy (with a coefficient of -28.3) and highly significant (with a *t*-value of -7.12).

Interpreting this is not straight-forward. If this reflects large R&D expenditures and low income for start-up firms, it should be captured by the age-variable. We would expect that getting the subsidy as a tax cut or in cash doesn't matter to the firms. But it may be that broke firms have larger utility of liquidity, and thus get an extra incentive to write applications. Also, it may be the case that some managers/firms are just good at getting the best from the public sector, both tax exemption and R&D subsidies.

## Problem 8

Assumption (3) states that the dependent variables, and also the regressors, should be i.i.d. This implies that the error terms are uncorrelated:  $cov(u_i, u_j) = 0, \quad i \neq j$ . This will likely not be the case when firms appear twice. It seems likely that firm characteristics are persistent over time, and that a firm with a large positive (negative) residual in 2003, will also have a positive (negative) residual in 2004. This may reflect that the firm for example has a large R&D department (compared to other firms with similar observable characteristics), which is likely to be persistent between years. If we use  $u_{i,t}$  to denote the residual of firm  $i$  in year  $t$ , this implies  $cov(u_{i,2003}, u_{i,2004}) \neq 0$ .

This is usually referred to as *autocorrelation*, and has a impact on the estimates similar to that of heteroskedasticity: The coefficient estimates will not be affected, but the estimated standard errors will. Thus, we will likely overstate the precision of the estimates, and may reject hypotheses we shouldn't have rejected. In Stata, we can control for such error term correlations using the option `cluster(orgnr)` to `regress`.

## Problem 9

The fact that only about 14 percent of the firms got a subsidy means that the subsidy cannot be anywhere near normally distributed. (The distribution of subsidies  $> 0$  is also somewhat skewed, that is a minor problem however.) Thus, the regression model we have used so far may be inappropriate. In the appendix I have included output from a regression using instead  $y$ , a dummy for whether the firm got any subsidy, as the dependent variable.<sup>15</sup>

It's difficult to make meaningful comparisons of the magnitudes of the coefficients, given the different natures of  $y$  (binary) and  $RD\_subsidy$  (continuous). However, comparing the current regression output with that from problem 6, we see that all variables have the same signs, and all the  $t$ -values are similar. Thus the qualitative picture stays the same.

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<sup>15</sup>The standard approach when using a binary left hand side variable is to use either a **logit** or **probit** model. The linear model we use, often called the linear probability model, has some conceptual and practical problems, but is still consistent and often considered a good starting point.



## Appendix: Stata log

```

-----
log:  \\Balder\540$\kir\Internett\Annet\ECON 4135\wp2.log
log type:  text
opened on:  25 Oct 2007, 15:20:57

. /* Stata-code for written paper II
> * ECON 4135 , Autumn 2007 */
. . . * Problem 1 . use manuf2003,clear

. su emply- empl5

      Variable |      Obs      Mean   Std. Dev.      Min      Max
-----+-----
      emply |      4215   30.32076   160.9113         0    3378
    share_high |      4215   .0273272   .1117486         0         1
      VA_empl |      4084   397.3964   758.7823   -13000    28340
taxposition |      4215   .5333333   .4989468         0         1
    firmage_10y |      4215   .462159   .4986252         0         1
-----+-----
    RD_subsidy |      4215   58.72997   182.738         0    1600
      y |      4215   .1399763   .3470036         0         1
      empl2 |      4215   .2185053   .4132811         0         1
      empl3 |      4215   .2994069   .4580526         0         1
      empl4 |      4215   .0483986   .2146324         0         1
-----+-----
      empl5 |      4215   .0455516   .2085353         0         1

. su emply- empl5 if y==1

      Variable |      Obs      Mean   Std. Dev.      Min      Max
-----+-----
      emply |      590   60.76271   191.3244         0    3378
    share_high |      590   .0570886   .1306425         0         1
      VA_empl |      583   424.6133   638.6254   -3777    11049
taxposition |      590   .4508475   .4980004         0         1
    firmage_10y |      590   .4457627   .4974714         0         1
-----+-----
    RD_subsidy |      590   419.5709   295.3912     2.515    1600
      y |      590         1         0         1         1
      empl2 |      590   .1576271   .3647002         0         1
      empl3 |      590   .4389831   .4966841         0         1
      empl4 |      590   .1440678   .3514564         0         1
-----+-----
      empl5 |      590   .1118644   .315467         0         1

```

```
. su RD_subsidy if y==1,de
```

RD_subsidy				
Percentiles		Smallest		
1%	12.84	2.515		
5%	38.122	7.564		
10%	78.458	9.153	Obs	590
25%	165.002	10.781	Sum of Wgt.	590
50%	352.2135		Mean	419.5709
		Largest	Std. Dev.	295.3912
75%	720	1440		
90%	800	1600	Variance	87255.96
95%	822.551	1600	Skewness	.6600611
99%	1223.629	1600	Kurtosis	3.115796

```
. . * Problem 2 . reg RD tax share VA firm emply
```

Source	SS	df	MS	Number of obs =	4084
Model	5176058.31	5	1035211.66	F( 5, 4078) =	31.65
Residual	133374640	4078	32705.8951	Prob > F =	0.0000
Total	138550699	4083	33933.5534	R-squared =	0.0374
				Adj R-squared =	0.0362
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RD_subsidy	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
taxposition	-25.62379	5.72527	-4.48	0.000	-36.84845	-14.39914
share_high	220.8335	25.34942	8.71	0.000	171.1348	270.5322
VA_empl	.0027261	.0037569	0.73	0.468	-.0046394	.0100916
firmage_10y	-8.516276	5.701123	-1.49	0.135	-19.69359	2.661037
emply	.1238319	.0173638	7.13	0.000	.0897894	.1578744
_cons	66.48483	5.233204	12.70	0.000	56.2249	76.74477

```
. /* Note: We only need to use as many letters of variable names,
> * as to make them unique and thus understandable to Stata.
> * Also note, Stata is case-sensitive! */
. . use manuf2004
```

```
. reg RD tax share VA firm emply
```

Source	SS	df	MS	Number of obs =	4100
Model	8744545.34	5	1748909.07	F( 5, 4094) =	50.34
				Prob > F =	0.0000

Residual		142226879	4094	34740.3222	R-squared	=	0.0579
-----+							
Total		150971424	4099	36831.2819	Adj R-squared	=	0.0568
					Root MSE	=	186.39

RD_subsidy		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+						
taxposition		-41.80898	6.0566	-6.90	0.000	-53.68321 -29.93475
share_high		280.488	26.13402	10.73	0.000	229.2511 331.7249
VA_empl		.0111815	.0028806	3.88	0.000	.0055339 .0168291
firmage_10y		6.076629	5.837064	1.04	0.298	-5.367189 17.52045
emply		.1288355	.017327	7.44	0.000	.0948652 .1628059
_cons		73.28736	5.735425	12.78	0.000	62.04281 84.53191

```

. . use manuf2003

. reg RD tax share VA firm emply ,robust

```

Linear regression		Number of
obs = 4084		
		F( 5, 4078) = 12.15
		Prob > F = 0.0000
		R-squared = 0.0374
		Root MSE = 180.85

RD_subsidy		Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
-----+						
taxposition		-25.62379	5.712823	-4.49	0.000	-36.82404 -14.42354
share_high		220.8335	41.06181	5.38	0.000	140.33 301.3371
VA_empl		.0027261	.0053747	0.51	0.612	-.0078113 .0132635
firmage_10y		-8.516276	5.637936	-1.51	0.131	-19.56971 2.537156
emply		.1238319	.0444131	2.79	0.005	.0367579 .2109059
_cons		66.48483	5.617412	11.84	0.000	55.47164 77.49803

```

. use manuf2004

. reg RD tax share VA firm emply ,robust

```

Linear regression		Number of
obs = 4100		
		F( 5, 4094) = 17.77
		Prob > F = 0.0000
		R-squared = 0.0579

Root MSE = 186.39

```
-----+-----
```

RD_subsidy	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
taxposition	-41.80898	6.433852	-6.50	0.000	-54.42283	-29.19514
share_high	280.488	47.55889	5.90	0.000	187.2467	373.7292
VA_empl	.0111815	.0047637	2.35	0.019	.0018421	.0205209
firmage_10y	6.076629	5.819209	1.04	0.296	-5.332184	17.48544
empl	.1288355	.0458719	2.81	0.005	.0389017	.2187694
_cons	73.28736	6.080932	12.05	0.000	61.36543	85.20929

```
-----+-----
```

. . \* Problem 3 . use manuf2003

. reg RD tax share VA firm empl2-empl5

Source	SS	df	MS	Number of obs =	4084
Model	16025704.5	8	2003213.06	F( 8, 4075) =	66.62
Residual	122524994	4075	30067.4832	Prob > F =	0.0000
				R-squared =	0.1157
				Adj R-squared =	0.1139
Total	138550699	4083	33933.5534	Root MSE =	173.4

```
-----+-----
```

RD_subsidy	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
taxposition	-20.63591	5.498826	-3.75	0.000	-31.41661	-9.855205
share_high	229.1315	24.36033	9.41	0.000	181.3719	276.891
VA_empl	-.0000786	.0036064	-0.02	0.983	-.007149	.0069918
firmage_10y	7.31752	5.552952	1.32	0.188	-3.569299	18.20434
empl2	22.09559	7.289712	3.03	0.002	7.803771	36.38741
empl3	75.84837	6.724101	11.28	0.000	62.66546	89.03128
empl4	187.8132	13.00368	14.44	0.000	162.3189	213.3076
empl5	189.2877	13.37939	14.15	0.000	163.0568	215.5186
_cons	14.61535	6.425678	2.27	0.023	2.017514	27.21319

```
-----+-----
```

. . \* Problem 4

. di "VA\_empl, 99% ci lower bound: " %9.5g \_b[VA] - invttail(e(df\_r),0.005)\*\_se[VA]  
 VA\_empl, 99% ci lower bound: -.0093723

. di "VA\_empl, 99% ci upper bound: " %9.5g \_b[VA] + invttail(e(df\_r),0.0055)\*\_se[VA]  
 VA\_empl, 99% ci upper bound: .0092151

. lincom VA ,level(99)

( 1) VA\_empl = 0

RD_subsidy	Coef.	Std. Err.	t	P> t	[99% Conf. Interval]	
(1)	-.0000786	.0036064	-0.02	0.983	-.0093723	.0092151

. reg RD tax share VA firm empl2-empl5 ,level(99)

Source	SS	df	MS	Number of obs = 4084		
Model	16025704.5	8	2003213.06	F( 8, 4075) = 66.62		
Residual	122524994	4075	30067.4832	Prob > F = 0.0000		
Total	138550699	4083	33933.5534	R-squared = 0.1157		
				Adj R-squared = 0.1139		
				Root MSE = 173.4		

RD_subsidy	Coef.	Std. Err.	t	P> t	[99% Conf. Interval]	
taxposition	-20.63591	5.498826	-3.75	0.000	-34.80658	-6.465234
share_high	229.1315	24.36033	9.41	0.000	166.354	291.9089
VA_empl	-.0000786	.0036064	-0.02	0.983	-.0093723	.0092151
firmage_10y	7.31752	5.552952	1.32	0.188	-6.992639	21.62768
empl2	22.09559	7.289712	3.03	0.002	3.309736	40.88144
empl3	75.84837	6.724101	11.28	0.000	58.52012	93.17662
empl4	187.8132	13.00368	14.44	0.000	154.3023	221.3242
empl5	189.2877	13.37939	14.15	0.000	154.8085	223.7669
_cons	14.61535	6.425678	2.27	0.023	-1.943853	31.17456

. . \* Problem 5 . use manuf2004

. reg RD tax share VA firm empl2-empl5 ,level(99)

Source	SS	df	MS	Number of obs = 4100		
Model	22872312.7	8	2859039.09	F( 8, 4091) = 91.31		
Residual	128099112	4091	31312.4203	Prob > F = 0.0000		
Total	150971424	4099	36831.2819	R-squared = 0.1515		
				Adj R-squared = 0.1498		
				Root MSE = 176.95		

RD_subsidy	Coef.	Std. Err.	t	P> t	[99% Conf. Interval]	
taxposition	-36.38267	5.768243	-6.31	0.000	-51.24762	-21.51773

```

share_high | 290.8909 24.86771 11.70 0.000 226.8061 354.9758
VA_empl | .0101209 .0027357 3.70 0.000 .0030709 .0171709
firmage_10y | 23.0223 5.621733 4.10 0.000 8.534916 37.50969
empl2 | 24.42863 7.418571 3.29 0.001 5.310742 43.54653
empl3 | 79.16296 6.852952 11.55 0.000 61.50268 96.82323
empl4 | 223.5669 13.14094 17.01 0.000 189.7023 257.4315
empl5 | 206.6539 13.61469 15.18 0.000 171.5684 241.7394
_cons | 14.71667 6.856328 2.15 0.032 -2.952307 32.38564

```

```

-----
. di "test statistic: "
(-20.63591--36.38267)/sqrt(5.498826^2+5.768243^2) test statistic:
1.9759281

```

```

. . * Problem 6 . use manuf2003

```

```

. append using manuf2004

```

```

. gen d_2003 = year==2003

```

```

. reg RD tax share VA firm empl2-empl5 d_2003

```

Source	SS	df	MS	Number of obs =	8184
Model	38364878.2	9	4262764.25	F( 9, 8174) =	138.68
Residual	251259295	8174	30738.8421	Prob > F =	0.0000
				R-squared =	0.1325
				Adj R-squared =	0.1315
Total	289624173	8183	35393.3977	Root MSE =	175.32

RD_subsidy	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
taxposition	-28.34934	3.981638	-7.12	0.000	-36.15436 -20.54432
share_high	259.1364	17.40769	14.89	0.000	225.013 293.2599
VA_empl	.0063219	.0021743	2.91	0.004	.0020597 .0105841
firmage_10y	15.0352	3.953013	3.80	0.000	7.286293 22.78412
empl2	23.17059	5.204471	4.45	0.000	12.96851 33.37268
empl3	77.43843	4.80391	16.12	0.000	68.02155 86.85531
empl4	205.3759	9.250252	22.20	0.000	187.243 223.5087
empl5	198.003	9.549805	20.73	0.000	179.2829 216.723
d_2003	-8.403869	3.893134	-2.16	0.031	-16.0354 -.7723367
_cons	18.25918	5.192032	3.52	0.000	8.081481 28.43689

```

. . * Problem 9 . reg y tax share VA firm empl2-empl5 d_2003

```

Source	SS	df	MS	Number of obs =	8184
--------	----	----	----	-----------------	------

```

-----+-----
      Model | 123.907048      9 13.7674498
      Residual | 926.122155 8174 .113300973
-----+-----
      Total | 1050.0292 8183 .128318368

```

F( 9, 8174) = 121.51  
 Prob > F = 0.0000  
 R-squared = 0.1180  
 Adj R-squared = 0.1170  
 Root MSE = .3366

```

-----+-----
            y |      Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----
taxposition |  -.0523249   .0076442    -6.84  0.000   - .0673096   - .0373402
share_high  |   .395625   .0334206    11.84  0.000   .3301122   .4611379
VA_empl    |   .0000112  4.17e-06     2.69  0.007   3.06e-06   .0000194
firmage_10y |  .0258087   .0075893     3.40  0.001   .0109318   .0406857
empl2      |   .0553754   .0099919     5.54  0.000   .0357887   .0749621
empl3      |   .1772021   .0092229    19.21  0.000   .1591228   .1952813
empl4      |   .3975509   .0177593    22.39  0.000   .3627381   .4323637
empl5      |   .3072914   .0183344    16.76  0.000   .2713512   .3432315
d_2003     |  -.019209   .0074743    -2.57  0.010  - .0338606  - .0045574
_cons      |   .0614897   .0099681     6.17  0.000   .0419498   .0810296
-----+-----

```

. predict yhat (option xb assumed; fitted values) (235 missing values generated)

. su yhat ,de

Fitted values

```

-----+-----
      Percentiles   Smallest
1%   -.0071431     -.1471096
5%   .0111495     -.0884523
10%  .0192287     -.0681077   Obs           8184
25%  .0504399     -.0291871   Sum of Wgt.   8184

50%  .1189307
      Largest
75%  .2225205     .6552632
90%  .3330851     .6632593   Variance      .015142
95%  .4131974     .6661162   Skewness      .9536523
99%  .4793803     .6746188   Kurtosis      3.387368
-----+-----

```

. log close

log: \\Balder\540\$\kir\Internett\Annet\ECON 4135\wp2.log

log type: text

closed on: 25 Oct 2007, 15:20:59